Determining the Change in Retail Prices of Veal in Turkey by GARCH Method Between 2014-2017^[1]

Mehmet Saltuk ARIKAN ^{1,a} Mustafa Bahadır ÇEVRİMLİ ^{2,b} Ahmet Cumhur AKIN ^{3,c} Mustafa Agah TEKİNDAL ^{4,d}

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- ¹ Department of Animal Health Economics and Management, Firat University Faculty of Veterinary Medicine, TR-23119 Elazig - TURKEY
- ² Department of Livestock Economics and Management, Selcuk University Faculty of Veterinary Medicine, TR-42003 Konya - TURKEY
- ³ Department of Animal Health Economics and Management, Mehmet Akif Ersoy University Faculty of Veterinary Medicine, TR-15030 Burdur TURKEY
- ⁴ Department of Biostatistic, Selcuk University Faculty of Veterinary Medicine, TR-42003 Konya TURKEY
- ^a ORCID: 0000-0003-4862-1706; ^b ORCID: 0000-0001-5888-242X; ^c ORCID: 0000-0003-3732-0529; ^d ORCID: 0000-0002-4060-7048

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Abstract

In Turkey, the demand for veal in the red meat market is steadily increasing. In the formation of this demand and price; additional factors such as consumer preferences, government interventions in the red meat market, import decisions, also the implemented policies and subsidies are effective besides the main factors. However, the fact that such practices do not affect the veal prices in the market makes it impossible to ensure price stability and a sustainable market. The analysis of the change in retail prices is important for both producers and consumers so that the validity of this view can be tested. The aim of this study was to analyze the retail prices of minced meat and the veal cubes from the first week of January 2014 to the last week of December 2017 via using the volatility estimation modeling method in the context of the effectiveness of the policies and market interventions applied in the red meat sector implemented in Turkey. Non-stationary series were made stationary by differencing. By developing a Generalised Autoregressive Conditional Heteroskedasticity (GARCH 1,1) model using the series subject to a unit root test, it was found to be the optimal model that best explains the fluctuation of the prices. The results of the estimation show that the retail prices of veal cubes and minced meats fluctuated conspicuously in the said period and that the implemented policies and the market interventions were not adequate to eliminate the instability and uncertainty of the prices.

Keywords: GARCH, Price volatility, Time series, Turkey, Veal

Türkiye'de Perakende Dana Eti Fiyatlarındaki Değişimin Garch Yöntemiyle Belirlenmesi, 2014-2017

Öz

Türkiye'de kırmızı et tüketiminde büyükbaş hayvan etine olan talep her geçen gün artmaktadır. Bu talebin ve fiyatın oluşmasında tüketici tercihleri, kırmızı et piyasasına yapılan devlet müdahaleleri, alınan ithalat kararları, uygulanan politikalar ve desteklemeler de etkili olmaktadır. Ancak bu uygulamaların etkisinin reel piyasada dana eti fiyatlarına olumlu yansımaması sebebiyle, fiyat istikrarının sağlanması ve sürdürülebilir yapıya kavuşması imkânsızlaşmaktadır. Bu çalışmada; Ocak 2014 ilk hafta ile Aralık 2017 son haftası arasındaki döneme ait verilere göre Türkiye'de kırmızı et sektöründe uygulanan politikalar ve devlet müdahaleleri kapsamında perakende dana kıyma ve kuşbaşı ortalama fiyatları volatilite tahmin modelleme yöntemi ile analiz edilmesi amaçlanmıştır. Durağan olmayan serilerin farkları alınarak durağan hale getirilmiştir. Birim kök testi uygulanan serilere daha sonra Genelleştirilmiş Otoregresif Koşullu Değişen Varyans (GARCH 1,1) modeli kurularak, fiyatlardaki dalgalanmayı açıklayan en uygun model olduğu tespit edilmiştir. Tahmin sonuçları incelenen dönemde perakende dana kuşbaşı ve kıybaşı ve kıyma fiyatlarında belirgin dalgalanmaların yaşandığını göstermekle birlikte bu dönemde uygulanan politikaların ve piyasaya yapılan müdahalelerin fiyatlardaki istikrarısızlık ve belirsizliği gidermede yeterli olmadığını açıklamaktadır.

Anahtar sözcükler: GARCH, Fiyat oynaklığı, Zaman serisi, Türkiye, Dana eti

iletişim (Correspondence)

- +90 424 237 00 00-3995 Mobile: +90 505 3254651
- msarikan@firat.edu.tr

INTRODUCTION

In Turkey, the demand for red meat is met by sheep, goat, water buffalo and cattle producers, with beef constituting 87.67% of the total red meat production ^[1]. However, retail beef prices have substantially increased in recent years.

While retail veal prices rose by 130% between 2005-2013, retail mutton prices increased by 154% and the feed prices by 115%. The increases in the prices of red meat were higher than the increases in the CPI (Consumer Price Index), which was 118% in the same period ^[2]. This drastic increase in the retail prices resulted in approximately 30% decrease in the per capita consumption of red meat ^[3]. The increase in prices continued in the period from 2014 to 2017 as well, and retail prices of veal rose by 19.54% ^[4]. The increasing prices of red meat have been a subject of debate both in the media and among the policy-makers in Turkey ^[2].

The increasing prices both mean a higher level of uncertainty regarding the prices in the next years and poses a significant price risk, which adversely affects producer welfare where there are without risk protection mechanisms in place ^[5]. The increasing demand for red meat and global pressures on climate change have pushed up wholesale and retail prices of meat products in many parts of the world ^[6]. Abrupt changes in retail prices usually affect the food security for a part of the population, resulting in changes to the dietary habits ^[7,8].

Price volatility indicates the range in which the prices may vary in the future ^[9]. Knowledge of the degree of price volatility is very important in the sense that it would allow policy-makers to formulate appropriate strategies to protect against risks. In primary commodity markets such as the red meat market, production decisions are based on expected prices. When such expectations are not met, producers face the risk of decreasing revenues ^[10].

Price fluctuations affect all actors in the food supply chain. Volatility decreases the inputs of production and investments, causing a risk for producers who can react ^[11-13]. Unexpected price increases for consumers who spend a large portion of their income on food have a negative impact on sustainable food consumption ^[14].

It's a common fact that the volatilities in meat prices can be successfully modelled with the ARCH (Autoregressive Conditional Heteroscedasticity)/GARCH (Generalized Autoregressive Conditional Heteroscedasticity) time series models. GARCH is a model that can be used to analyse both mean and variance effects of policy and market structure changes, and in a changing environment such as the livestock market there is a need for such a model that combines the desirable aspects of dynamics, static equilibrium and price volatility ^[15,16]. It has been successfully used to analyse and explain the price changes of red meat in various countries ^[10,15,17,18]. Changes in prices of red meat products are a significant economic risk that affects the level of production. Thus, the research on the changes in retail prices of meat in Turkey is important to understand how to equilibrate the market.

This study is intended to analyse the average retail prices of minced meats and veal cubes under the government policies and interventions in the red meat sector of Turkey between 2014 and 2017 using the volatility estimation modelling method. The main purpose of this study is contributing the literature on the effectiveness of the policies and market interventions applied in the red meat sector. In the long run, this study also aims to provide significant information to the decision and policy makers about the econometric method that they can need.

MATERIAL and METHODS

Data Set

In order to explore the changes in meat prices, the weekly retail prices of minced and veal cubes in Turkey from the first week of 2014 to the last week of 2017 were used as the data for the analysis. The short-term asymmetric effect was determined by considering the time interval to measure the response of price changes. The data sets of the minced meats and veal cubes prices used in the analysis were obtained from the weekly bulletin of the Meat and Milk Board^[4].

Analysis Method

The analyses conducted using ARIMA model and ARCH/ GARCH models are intended to measure the index uncertainty. In the present study, due to the absence of future predictions, volatility of the price change series has been measured over the current asymmetric effect GARCH model instead of hybrid models. In order to measure the index uncertainty, the ARCH model suggested by Engle ^[19] for the first time and the GARCH model, which had been developed by Bollerslev ^[20] were used in the study.

The functioning of the ARCH model can be addressed as follows ^[21]:

Let's assume a regression model with k explanatory variables.

$$Y_{t} = \beta_{1} + \beta_{2} X_{2t} + \dots + \beta_{k} X_{kt} + \varepsilon_{t}$$
[1]

In equation (1), under the assumption that conditional information can be obtained at time (t-1), the error term ϵ_{tr} is normally distributed with zero means and variance $(\alpha_0 + \alpha_1 + \epsilon_{t-1}^2)$, which can be expressed as $\epsilon_t \sim N[0,(\alpha_0 + \alpha_1 + \epsilon_{t-1}^2)]$. While the error term with zero mean is one of the assumptions of the method of least squares, the addressing of the error term's variance at time t as a function of the square of the error term at t-1 was introduced by the ARCH model. Additionally, the expression of the error term's variance as $(\alpha_0 + \alpha_1 + \epsilon_{t-1}^2)$ is called the ARCH(1) process.

Since relatively long lags are used and a fixed lag structure is suggested in applying the ARCH model, some constraints are imposed on the parameters in the conditional variance equation. In an attempt to avoid problems with negative variance parameter estimates, Bollerslev ^[20] extended the ARCH model and developed a new one that allowed for both more past information and a more flexible lag structure. As the said model was developed not as an alternative to the ARCH model, but as a model intended to improve it, it is called generalised ARCH, or shortly GARCH ^[22].

Using the conditional variance equation, standard GARCH(p,q) process can be expressed as follows:

$$\sigma_{t}^{2} = \omega + \alpha \varepsilon_{t-1}^{2} + \beta \sigma_{t-1}^{2}$$
^[2]

The equation (2) indicating the conditional variance is a function of the mean (ω), ARCH term (ϵ^{2}_{t-1}) and GARCH term (σ^{2}_{t-1}). Therefore, p and q in GARCH(p, q) denote the ARCH term and GARCH term, respectively. In the ARCH(q) process, conditional variance is a function of only the past sample variances, whereas the GARCH(p, q) process allows lagged conditional variances to enter as well.

When working with time series, one should always pay attention to the stationarity of the data. Time series data of economic and financial variables usually involve trend and seasonality, which may violate the principle of stationary time series ^[23]. A stationary time series can be defined as mean and variance are independent over time. Where time series are not stationary, the estimated econometric models may produce misleading results. Thus, in the econometric analyses conducted using time series, unit root test (stationarity test) is performed on the time series. Therefore, Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) tests were conducted to test whether the time series had a unit root (stationarity). Results of the unit root test are given in *Fig. 1* and *Fig. 2* as well as in *Table 1* and *Table 2*.

Fig. 1 and *Fig.* 2 present the change in the prices of veal cubes and minced meats over time. It is clear that the variables do not follow a stationary course during the periods under consideration. The veal cubes and minced meats series were rendered stationary after the ADF and PP tests.

RESULTS

Decreases and increases in the retail prices of minced meats and veal cubes indicate a significant level of price fluctuation. Following the increases in October 2014 and May 2015, excessive increases are observed in June 2017 (*Fig. 1, Fig. 2*).

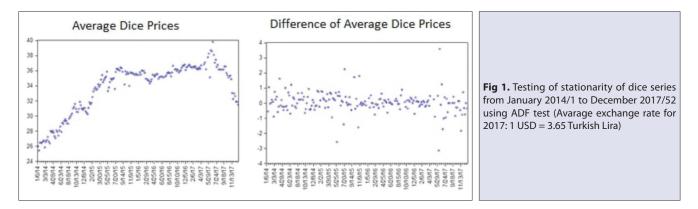
The series used in the model was compiled using the weekly data from January 2014/1 to December 2017/52 and the ARCH and GARCH methods. Firstly, the nonstationary series was made stationary by taking its logarithm and first difference (Fig. 1, Fig. 2). Secondly, it was found that inflation followed the AR(1) and MA(1) process by looking at the correlogram of ACF and PACF (Table 3). The modelling with AR(1) and MA(1) showed that both were significant. AIC and BIC values of the results obtained by running AR(1) and MA(1) individually were compared. As AIC and BIC values of AR(1) were lower than those of MA(1), AR(1) model was selected (Table 4, Table 5). Finally, the estimated ARCH/GARCH model was solved using the maximum likelihood (ML) technique, as shown in Table 4 and Table 5, and it was found that the model had the effect of GARCH (1, 1) and no effect of a higher degree of GARCH. The presence of the GARCH(1, 1) effect in the index series indicates that the GARCH variance series obtained as a result of the solution can be used as a measure of uncertainty in the veal cubes and minced meats series from January 2014/1 to December 2017/52.

The first difference of the data was taken to make the veal cubes series in *Fig. 1* stationary. The series became stationary after taking its first difference.

The first difference of the data was taken to make the minced meats series in *Fig. 2* stationary. The series became stationary after taking its first difference.

ADF and PP unit root tests indicate that the veal cubes series from January 2014/1 to December 2017/52 became stationary after taking its first difference.

ADF and PP test values of the veal cubes series from January 2014/1 to December 2017/52 after taking its first difference are significant at a significance level of 1%, 5%



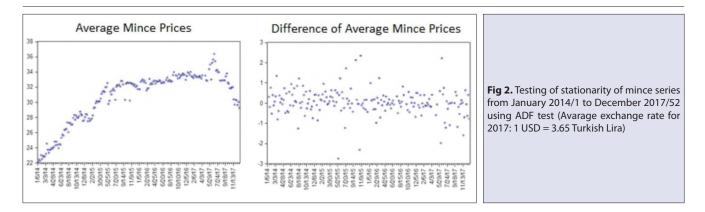


Table 1. Unit root tests coefficients for the dice series						
Variable	Level Value	First Difference				
Veal cubes	AI	DF	РР			
	-12.01921 (12)	-3.467851 (12)	-3.471454 (12)	-2.879494 (12)		

Table 2. Unit root tests coefficients for the mince series						
Variable	Level Value	First Difference	Level Value	First Difference		
Minced meats	AI	DF	РР			
	18.09752 (12)	-3.466176 (12)	-0.145169 (12)	-3.470934 (12)		

and 10%, respectively. The values in the parentheses are the optimal lag lengths.

ADF and PP unit root tests indicate that the mince series from January 2014/1 to December 2017/52 became stationary after taking its first difference.

ADF and PP test values of the mince series from January 2014/1 to December 2017/52 after taking its first difference are significant at a significance level of 1%, 5% and 10%, respectively. The values in the parentheses are the optimal lag lengths.

In this framework, the model developed below for the mince series from January 2014/1 to December 2017/52 is significant.

$$\sigma_t^2 = 0.087583 + 0.154908\sigma_{t-1}^2 - 0.647365\varepsilon_{t-1}^2$$
[3]

The model indicates that the mince price is affected by the price in the previous week, with 1 unit of increase in the mince price being followed by a 0.15 unit of increase in the mince price in the next week.

$$\sigma_t^2 = 0.140914 + 0.787804\sigma_{t-1}^2 - 0.133606\varepsilon_{t-1}^2 \quad [4]$$

The model indicates that the veal cubes price is affected by the price in the previous week, with 1 unit of increase in the veal cubes price being followed by a 0.78 unit of increase in the veal cubes price in the next week.

As an econometric result, it was found that the minced meats and veal cubes series from January 2014/1 to

December 2017/52 between 06.01.2014 and 25.12.2017 had a quite varying variance.

In the models set on the prices of minced meats and veal cubes, RESID (-1) parameter shows the effect of shocks on price changes and GARCH (-1) parameter shows the effect of the previous period volatility on the current period volatility. The fact that the sum of both parameters is close to one means that it will have a long time effect on price volatility if a possible equilibrium situation deteriorates.

According to the results of the analysis, the veal cubes are more sensitive to the price changes compared to minced meats.

DISCUSSION

This study analysed the volatility of the retail prices of minced meats and veal cubes using volatility estimation modelling for the period between 2014 and 2017, and the results of the GARCH(1, 1) model show that it is the optimal model that best explains the change in the prices.

Analysis results indicate that the retail prices of minced meats and veal cubes saw excessive volatility in the period under consideration. Various policies and government interventions were implemented in the red meat sector in an effort to decrease the rising meat prices. In order to prevent price increases in October and November 2014, the General Directorate of Meat and Milk Board was authorised to import 15 thousand tons of beef with zero custom duty ^[24], and the custom duty for livestock imports

Transformer	Manaha	Q Statistic of Error Terms			_		Q Statistic of	Error Terms		
Type of Meat Months	Months	AC	PAC	Q-Stat	Prob	Type of Meat	AC	PAC	Q-Stat	Prol
	1	-0.264	-0.264	14.059	0.000		-0.283	-0.283	16.148	0.00
	2	-0.094	-0.176	15.839	0.000		-0.066	-0.159	17.027	0.00
	3	-0.039	-0.127	16.150	0.001		-0.012	-0.087	17.054	0.00
	4	0.027	-0.045	16.304	0.003		-0.020	-0.069	17.135	0.00
	5	0.142	0.132	20.429	0.001		0.216	0.204	26.746	0.00
	6	-0.076	0.009	21.612	0.001		-0.106	0.023	29.069	0.00
	7	0.035	0.066	21.870	0.003		0.077	0.117	30.313	0.0
	8	-0.110	-0.084	24.412	0.002		-0.201	-0.168	38.709	0.0
	9	0.185	0.146	31.612	0.000		0.159	0.078	44.020	0.0
	10	-0.063	-0.009	32.447	0.000		-0.015	-0.045	44.070	0.0
	11	-0.010	0.017	32.466	0.001		0.044	0.093	44.489	0.0
	12	0.007	0.007	32.477	0.001		0.009	0.015	44.508	0.0
	13	0.031	0.061	32.682	0.002		-0.059	0.062	45.248	0.0
	14	0.023	0.002	32.800	0.003		0.073	0.002	46.392	0.0
	15	-0.085	-0.047	34.380	0.003	Veal cubes	-0.059	-0.001	47.142	0.0
	16	0.153	0.120	39.494	0.001		0.113	0.032	49.903	0.0
	17	-0.054	0.036	40.126	0.001		-0.056	0.027	50.583	0.0
	18	0.034	0.028	40.379	0.002		0.041	0.039	50.945	0.0
linced meats	19	-0.120	-0.103	43.591	0.001		-0.002	0.036	50.946	0.0
	20	0.057	0.010	44.310	0.001		-0.044	-0.032	51.386	0.0
	21	0.102	0.066	46.649	0.001		0.076	0.023	52.690	0.0
	22	-0.082	-0.039	48.164	0.001		-0.013	0.024	52.728	0.0
	23	0.089	0.081	49.950	0.001		0.066	0.061	53.722	0.0
	24	-0.091	0.013	51.852	0.001		-0.055	0.015	54.410	0.0
	25	0.039	-0.030	52.204	0.001		0.005	0.001	54.415	0.0
	26	0.004	-0.011	52.208	0.002		0.025	0.011	54.563	0.0
	27	0.008	-0.003	52.222	0.002		-0.049	-0.069	55.118	0.0
	28	-0.001	0.019	52.223	0.004		0.034	-0.038	55.394	0.0
	29	-0.008	-0.001	52.239	0.005		0.006	0.025	55.402	0.0
	30	0.014	-0.030	52.288	0.007		0.019	0.028	55.491	0.0
-	31	-0.048	-0.015	52.840	0.009		-0.071	-0.047	56.681	0.0
	32	0.065	0.009	53.859	0.009		0.002	-0.043	56.682	0.0
	33	0.033	0.070	54.127	0.012		0.089	0.062	58.574	0.0
	34	0.008	0.040	54.143	0.016		-0.018	0.013	58.654	0.0
	35	-0.026	0.046	54.312	0.020		-0.041	-0.055	59.056	0.0
	36	-0.047	-0.052	54.843	0.023		-0.041	-0.039	59.471	0.00

by the private sector was reduced to 15%, provided that imported livestock would be used for stockbreeding ^[25]. Although the decisions taken were effective in reducing the retail prices of minced meats and veal cubes until the end of 2014, the prices started to elevate again as from early 2015 until May. The rising prices adversely affected the prices of foods and food services in this period ^[26]. In an attempt to prevent the price increases, first the custom duty for imports was reduced to 0%, and then the Meat and Milk Board was authorised to import 30,000 tons of fresh or frozen bovine meat ^[27,28]. To prevent the increase in red meat prices which were rising due to the lack of sufficient cattle for slaughter in Turkey, a market intervention was implemented through the Meat and Milk Board in March 2017, supplying butchers with fresh beef carcass and food and meat industrialists with frozen beef carcass through importation ^[29]. The upward movement in red meat prices observed since February 2017 continued in June, and the increase in prices was reported to be basically arising from the shortage of domestic supply ^[30]. Pursuant to the decision taken in response to the drastic price increase in June, the custom

	analysis of the mince series from January 2014/1 to Dec Coefficient Std. Error				
Variable	Variance	Equation	z-Statistic	Р	
С	0.087583	0.027537	3.180620	0.0015	
RESID(-1)^2	0.154908	0.058710	2.638520	0.0083	
GARCH(-1)	0.647365	0.094256	6.868192	0.0001	

Table 5. The result of GARCH analysis of the dice series from January 2014/1 to December 2017/52							
Veriable	Coefficient Std. Error		z-Statistic				
Variable	Variance	Equation	z-statistic	р			
С	0.140914	0.019868	7.092685	0.0001			
RESID(-1)^2	0.787804	0.152150	5.177816	0.0001			
GARCH(-1)	0.133606	0.052468	2.546417	0.0109			
GARCH (1, 1): Conditional Variance Equation; P=0.001							

duty for bovine animals was reduced from 135% to 26%, and the custom duty for bovine meat ranging between 100-225% was reduced to 40% [31]. This decision granted the authorisation to import livestock and carcass, which had been exclusively held by the Meat and Milk Board, to the private sector as well. One month after the issuance of this decree, the Meat and Milk Board was granted authorisation to import 90 thousand tons of red meat and 975 thousand head of livestock, 500 thousand head being bovine and 475 thousand head being ovine [32].

Price uncertainty seems to be a very significant restrictive factor for meat production where risk protection mechanisms do not exist. As it affects the production decisions, it restricts the efforts of businesses to build new capacity, invest in advanced technologies and enter new markets ^[10]. The Turkish livestock sector is subject to excessive price fluctuations, low profitability and unforeseeable intervention policies [33].

Another factor leading to the price increases is the role of market concentration and market power, particularly at the level of retail prices^[2]. While the four-firm concentration (CR₄) in the red meat market of Turkey was 41% in 2007, it rose to 50% in 2012 ^[34,35]. The long marketing chain in the red meat market diminishes the effectiveness of government subsidies granted to producers and causes the producer share in consumer prices of many products to gradually decrease. This indicates that the subsidies granted are actually not transferred to producers, but rather to the processes after the production ^[36]. Due to the asymmetrical price transfer in the red meat market, there are no large-scale enterprises producer cooperatives that have the potential to become suppliers^[2]. As producers are small-scale enterprises, it is highly essential that they organise against large retailers and wholesalers so that they can have sufficient bargaining power under the current market conditions^[3].

The models developed for the time series data of the minced meats and veal cubes prices in the research show that increases in prices affect the next week's prices. The studies conducted on the same subject found that volatility of beef prices reacted heavily to the market movements ^[10], that a negative shock in the prices caused a higher volatility than a positive shock of the same degree [37], and that the fluctuations in beef prices resulted in fluctuations in the price of mutton and chicken [18].

Veal meat prices in Turkey were determined to be strictly affected from the economic and political instability, food crisis, natural events etc., and the prices of veal meat could return to normal values after a long time compared to the average price [38]. On the other hand, the dependence on beef and veal red meat has increased over the years and it reflected to the importation as well ^[39]. Red meat consumption has been sensitive to the price changes and always responded negatively to the beef meat pricing crisis ^[40]. Red meat imports, unfortunately, couldn't help the price increases and the market price of red meat was increased later because of a lower local supply [41]. On the other hand, due to the high level intermediary margins of beef meat marketing, red meat prices is increasing in Turkey^[42].

In conclusion, the results of the estimation prove that the retail prices of veal cubes and minced meats fluctuated conspicuously in the period under consideration and that the implemented policies and the market interventions were not adequate to eliminate the instability and uncertainty of the prices. The presence of oligopsonistic market conditions, the long marketing chain in the sector and the import dependency to meet the demand for red meat attenuate the market power. Hence, it is essential that the price uncertainty is reduced from the perspective of producers and concrete steps other than importation be taken to prevent increases in retail prices from the perspective of consumers so that the sector can grow and reach a level where it can compete globally.

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